Coronary Risk Estimation in Spain Using a Calibrated Framingham Function

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Introduction and objectives. The Framingham coronary heart disease (CHD) functions overestimate the risk of CHD in countries with a low incidence. Consequently, these functions should be calibrated for the purpose of primary prevention. Calibrated Framingham function charts of overall CHD risk for the Spanish population are presented.

Patients and methods. The Framingham functions were calibrated by substituting the prevalence of CHD risk factors and incidence found in Framingham with the same values for Spain. The Framingham function that included high-density lipoprotein (HDL) cholesterol was used. The 10-year probability of developing a CHD event was estimated for several combinations of risk factors and HDL levels ranging from 35 to 59 mg/dl. Color-coded charts were prepared that show the exact probability of CHD corresponding to each combination of risk factors, shown in separate cells on the chart.

Results. The event rate and prevalence of CHD risk factors differed considerably between Girona and Framingham. HDL < 35 mg/dl increased risk by approximately 50% and HDL > 60 mg/dl reduced it by 50%. The proportion of cells in which the 10-year probability of developing a CHD event was > 9% was 2.3 times higher and that of cells with a probability > 19% was 13 times lower in the chart calibrated for Spain than in the original Framingham charts.

Conclusions. The calibrated Framingham function may help to more accurately estimate the overall risk of CHD in the Spanish population for primary prevention purposes. The calibrated function should be validated, and the development of functions for the Spanish population should be promoted.

Key words: Coronary disease risk. Coronary artery disease. Coronary heart disease risk functions.

Estimación del riesgo coronario en España mediante la ecuación de Framingham calibrada

Introducción y objetivos. Las ecuaciones de Framingham sobrestiman el riesgo de enfermedad coronaria en los países cuya incidencia es baja. En éstos, la ecuación debería adaptarse para la correcta prevención de la enfermedad coronaria. Se presentan las tablas de riesgo coronario global de Framingham calibradas para la población española.

Pacientes y método. Se utilizó el procedimiento de calibración de la ecuación de Framingham, consistente en sustituir la prevalencia de factores de riesgo cardiovascular y la tasa de incidencia de acontecimientos coronarios de Framingham por las de nuestro medio. Se ha usado la ecuación de Framingham, que incluye el colesterol unido a lipoproteínas de alta densidad (cHDL). Se han calculado las probabilidades de acontecimiento a los 10 años y se han elaborado unas tablas con códigos de color y la probabilidad exacta en cada casilla correspondiente a las distintas combinaciones de los factores de riesgo clásicos, para una concentración de cHDL de 35-59 mg/dl.

Resultados. Las tasas de acontecimientos coronarios y la prevalencia de factores de riesgo difieren considerablemente entre la población estudiada y Framingham. Valores de cHDL < 35 mg/dl incrementan el riesgo en un 50% y los > 60 mg/dl lo reducen en un 50%, aproximadamente. La proporción de casillas con una probabilidad de acontecimiento coronario a los 10 años superior al 9% es 2,3 veces menor, y la de casillas con una probabilidad > 19% es 13 veces menor en las tablas calibradas que en las originales de Framingham.

Conclusiones. La función de Framingham calibrada puede constituir un instrumento para estimar con más precisión el riesgo coronario global en la prevención primaria de esta enfermedad en España. Su uso debe acompañarse de una validación apropiada y se debe trabajar en la elaboración de ecuaciones propias españolas.

INTRODUCTION

As much as two thirds of over the 40 000 deaths caused by ischemic heart disease in Spain every year, occur before the patient reaches the hospital. Optimal primary prevention could prevent some of these deaths, since it would lower the incidence of ischemic heart disease which has stabilized over the past 15 years. Estimations of chronic diseases projected for the year 2020 indicate that ischemic heart disease will continue to be the first cause of death in developed countries, and the third cause in underdeveloped countries. In this context, all efforts aimed at improving prevention of coronary heart disease are justified.

Estimating the probability of developing a coronary event based on individual risk factors is a valuable tool for attaining this target. Coronary disease prevention based on total risk evaluation allows for a better decision-making than any other approach to isolated components. Risk factors do not show a homogeneous distribution and their effects are not identical in all populations. Particularly, paradoxical situations have been described in Southern Europe, where a low incidence of ischemic heart disease and of its associated mortality holds within a context of high prevalence of risk factors or saturated fats consumption.

The Framingham cohort has contributed to clarify the causes of ischemic heart disease and other cardiovascular diseases. Its results have allowed to develop mathematical functions for measuring the individual risk of coronary events under existing risk factors. Coronary risk calculation based on the Framingham study is the procedure recommended nowadays. In Spain, predictive models with population data have not been developed yet. There is also evidence that Framingham equation overestimates coronary risk in some of the populations.

As population cohort studies are unavailable, some techniques were proposed to calibrate the Framingham coronary hazard function adjusted to local prevalence of cardiovascular risk factors and coronary event incidence. A calibrated function for our environment was created with this approach. The good concordance between hard coronary events predicted (symptomatic AMI, fatal or non-fatal) with this calibration and the actual number was confirmed by the myocardial infarction population registry of Girona REGICOR (Registre Gironí del Cor).

The aim of this study is to present tables of total coronary risk estimation for the following 10 years based on a calibrated Framingham function adapted to the Spanish population.

PATIENTS AND METHOD

The classical Framingham function includes all fatal or non-fatal coronary events, such as angina and asymptomatic Q-wave AMI detected by electrocardiography. There is also an adjusted equation for hard coronary events: symptomatic AMI, fatal or non-fatal.

The Framingham equation calibration for hard coronary events, the way it overestimates these events, and the advantages of calibration in our environment have already been described in detail. The Framingham function estimates the risk of coronary events to 10 years by comparison of risk of an evaluated individual with the population averages. Calculation is based on median age and risk factor prevalence considered for that population. The population average rate of events to 10 years is also included. In Spain, none of the results match those of the Framingham cohort. In the calibrated Framingham function, the average comparison element is replaced by a local element. To make this possible, a reliable estimation of local risk factor prevalence, as well as the local coronary event incidence rate considered, and the original equation coefficients, are needed. Calculation of the calibrated function is described in detail in the Appendix.

Tables with the general model of all the Framingham coronary events were calculated using the equation published in 1998 by Wilson et al, that follows the described calibration method. Wilson’s function included cHDL and degrees of hypertension based on systolic (SD) and diastolic (DBP) blood pressure. In our case, we used the risk factor prevalence and population data included in the REGICOR study.

The hard event incidence rates were obtained from the REGICOR population registry. Since 1988, this registry investigates all suspected AMI cases in 6 districts of Girona, with a 35 to 74 year old population of around 200 000 inhabitants. As angina and silent AMI incidence rates for Girona are unknown, a Framingham similar proportion was assumed, and the hard event rate was proportionally increased. The following ratio was used:

\[
\frac{H_{0}(t)}{FramAll} = \frac{H_{0}(t)}{FramHard}
\]
where t is the follow-up time, 10 years in our case; H0(t)/FramAll is the coronary event rate including Framingham angina and silent AMI, and H0(t)/FramHard is the rate of symptomatic AMI, fatal or non-fatal. This quotient value is 1.400 for males and 1.910 for females. The hard event rate in males of Girona in the REGICOR registry (3.5%) is multiplied by 1.400 to obtain an estimated total coronary event rate (4.9%). This allows calculation of the following rate of population without events:

\[ 100 - 4.9\% = 95.1\% \]

A similar procedure is used in females: the hard event rate (1.1%) is multiplied by its quotient (1.910) to obtain the estimated total coronary event rate (2.2%). The rate of female population without events is as follows:

\[ 100 - 2.2\% = 97.8\% \]

The tables are calculated for average HDL-C in 35 to 74 year old males (47.5 mg/dL) of Girona. As in most guidelines, there is a coded color for hazard intensities in the risk factor combinations for diabetic and non-diabetic males and females separately. A hazard value, rounded to the higher nearest integer, appears in every combination checkbox of the risk factor degrees described in Table 1.

### RESULTS

The risk factor distributions for Girona and Framingham are shown in Table 1. Some categories of male and female risk factors are different in both distributions, and coronary event incidence rates also differ as for gender and between both regions.

Figure 1A shows a table of risk factor combinations of AMI, fatal or non-fatal, fatal AMI with or without symptoms, and angina in males. Figure 1B corresponds to diabetic males. Figure 2A illustrates risk in females and Figure 2B corresponds to diabetic females.

The resulting risk for HDL-C<35 mg/dL is approximately 50% higher than risk observed on the tables,

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**TABLE 1. Cox proportional hazards model coefficients of the Framingham function for all coronary events incidence to 10 years and prevalence of degrees of each risk factor in Framingham, USA, and Girona, Spain**

<table>
<thead>
<tr>
<th>Risk Factor</th>
<th>Cox Model Coefficients</th>
<th>Framingham Prevalence</th>
<th>Girona Prevalence</th>
<th>Cox Model Coefficients</th>
<th>Framingham Prevalence</th>
<th>Girona Prevalence</th>
</tr>
</thead>
<tbody>
<tr>
<td>Age median, years</td>
<td>0.0483</td>
<td>48.3</td>
<td>54.6</td>
<td>0.3377</td>
<td>49.6</td>
<td>54.2</td>
</tr>
<tr>
<td>Quadratic age</td>
<td>-0.6595</td>
<td>7.5%</td>
<td>5.9%</td>
<td>-0.2614</td>
<td>7.9%</td>
<td>4.9%</td>
</tr>
<tr>
<td>Total cholesterol (mg/dL)&lt;160</td>
<td></td>
<td>31.3%</td>
<td>20.7%</td>
<td>0</td>
<td>30.3%</td>
<td>11.8%</td>
</tr>
<tr>
<td>160-199</td>
<td>0.1769</td>
<td>39.0%</td>
<td>34.8%</td>
<td>0.2077</td>
<td>32.7%</td>
<td>35.3%</td>
</tr>
<tr>
<td>200-239</td>
<td>0.5054</td>
<td>16.5%</td>
<td>28.1%</td>
<td>0.2439</td>
<td>20.0%</td>
<td>24.6%</td>
</tr>
<tr>
<td>240-279</td>
<td>0.6571</td>
<td>5.7%</td>
<td>10.5%</td>
<td>0.5351</td>
<td>9.1%</td>
<td>13.4%</td>
</tr>
<tr>
<td>Total cholesterol (mg/dL)&lt;35</td>
<td>0.4974</td>
<td>19.2%</td>
<td>17.0%</td>
<td>0.8431</td>
<td>4.3%</td>
<td>3.8%</td>
</tr>
<tr>
<td>35-44</td>
<td>0.2431</td>
<td>35.7%</td>
<td>31.9%</td>
<td>0.3780</td>
<td>14.9%</td>
<td>17.0%</td>
</tr>
<tr>
<td>45-49</td>
<td>0.0222</td>
<td>5.5%</td>
<td>12.8%</td>
<td>0.1979</td>
<td>12.4%</td>
<td>12.3%</td>
</tr>
<tr>
<td>50-59</td>
<td>-0.0511</td>
<td>19.0%</td>
<td>20.9%</td>
<td>0</td>
<td>27.7%</td>
<td>26.0%</td>
</tr>
<tr>
<td>59-64</td>
<td>-0.4866</td>
<td>10.6%</td>
<td>17.5%</td>
<td>-0.4295</td>
<td>40.7%</td>
<td>40.9%</td>
</tr>
<tr>
<td>Blood pressure, mm Hg</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Optimal (SBP&lt;120)/DBP&lt;80</td>
<td>0.0023</td>
<td>20.2%</td>
<td>18.1%</td>
<td>-0.5336</td>
<td>34.8%</td>
<td>27.5%</td>
</tr>
<tr>
<td>Normal (SBP 120-129)/DBP 80-84</td>
<td>0</td>
<td>24.3%</td>
<td>19.3%</td>
<td>0</td>
<td>48.6%</td>
<td>16.5%</td>
</tr>
<tr>
<td>Normal high (SBP 130-139)/DBP 85-89</td>
<td>0.2832</td>
<td>20.2%</td>
<td>20.1%</td>
<td>-0.0677</td>
<td>15.0%</td>
<td>16.2%</td>
</tr>
<tr>
<td>Degree I (SBP 140-159)/DBP 90-99</td>
<td>0.5217</td>
<td>22.5%</td>
<td>28.4%</td>
<td>0.2629</td>
<td>18.6%</td>
<td>28.6%</td>
</tr>
<tr>
<td>Degree II-III (SBP≥160)/DBP≥100</td>
<td>0.6186</td>
<td>12.8%</td>
<td>14.1%</td>
<td>0.4657</td>
<td>10.0%</td>
<td>11.5%</td>
</tr>
<tr>
<td>Diabetes</td>
<td>0.4284</td>
<td>5.0%</td>
<td>9.3%</td>
<td>0.5963</td>
<td>3.8%</td>
<td>6.7%</td>
</tr>
<tr>
<td>Smoking</td>
<td>0.5234</td>
<td>40.3%</td>
<td>42.3%</td>
<td>0.2925</td>
<td>37.8%</td>
<td>12.0%</td>
</tr>
<tr>
<td>Incidence to 10 years</td>
<td></td>
<td>Of ischemic heart disease</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

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Sources:
- Cox proportional hazards model coefficients of the Framingham function for all coronary events incidence to 10 years and prevalence of degrees of each risk factor in Framingham, USA, and Girona, Spain.
whereas risk for HDL-C>59 mg/dL is approximately 50% lower. The risk factor combinations proportion that establishes a severe or very severe ischemic heart disease hazard (i.e., 19% risk to ten years) in the calibrated tables is 13 times less than in the original tables (Table 2). The proportion of factor combinations that lead from moderate to very high risk is 2.7 times lower. This result is very noticeable in non-diabetic males and females.

DISCUSSION

Some tables of total coronary risk, calibrated by including risk factor prevalence and estimated event rates in our environment, are presented for use in the Spanish population. The table values produce globally a percentage of risk factor combinations (checkboxes) with a moderate to very high total hazard that is 2.7 times less than in the original tables. In general, the resulting hazard for risk factor combinations is much lower.

Difficulties in extrapolating the Framingham function to certain low incidence regions have been described occasionally in the general population\textsuperscript{19-23,26,27} and in diabetics,\textsuperscript{24} although it matches reality in many parts of the world.\textsuperscript{29,30} The European coronary disease prevention guidelines expressly recommend that efforts are directed to adapt predictive functions to local requirements.\textsuperscript{5} This study follows such recommendation by applying a proven calibration method.\textsuperscript{5,19} Several approximations have been also used to adapt the equations.\textsuperscript{3,22} In the future, any other adaptations should include estimation of the maximum number of coronary events, and even of cardiovascular events.
Anyway, predicting exclusively fatal or hard events, now that ischemic heart disease mortality is falling and acute coronary syndromes are treated earlier with less complications, would render a biased and confusing vision in clinical practice.

The only cohort study with over 20 years follow-up, including a masculine working population only, demonstrated that observed coronary events depending on the degree of exposure to risk factors were less than expected when the study was initiated in 1968. Corrections obtained in our study match this finding, although according to the Manresa study it is not possible to measure the difference between values observed and values expected due to risk factor prevalence when follow-up started.

The authors of the Seven Countries and other previous studies, observed that smoking, systolic blood pressure, age and body mass index were the most important predictive variables for ischemic heart disease mortality to 6 years in Southern Europe (largest standardized coefficient in model), whereas cholesterol was not associated significantly with any coronary event risk.

Coronary risk calculation allows to establish the prevention targets for each patient individually. The effect of demonstrating a lower estimation, as is the aim of this study, could influence decision-making in pharmacological management of hypercholesterolemia. The current coronary heart disease prevention guidelines indicate that pharmacological treatment should be initiated in patients with a minimum 20% age risk to 10 years or projected to 60 years, if after management with life style measures during 6 months, total cholesterol is not below 190 mg/dL, and LDL-C is not
below 115 mg/dL, or 130 mg/dL in the Spanish recommendations. Pharmacological treatment for hypertension depends on the same degree of risk. Nevertheless, the following points explain why the present calibration should not be used as a justification for reducing the pharmacological treatment of hypertension:

a) negative clinical consequences of hypertension are many (stroke, mainly), for a near 50% attributable risk), and b) hypertension rates in Spain are highly unknown and scarcely controlled, as has been confirmed repeatedly.

Study characteristics and limitations

AMI incidence in Girona appears in the IBERICA (Investigación y Búsqueda Específica y Registro de Isquemia Coronaria Aguda) study as 15% below the Spanish average. External validation (applicability) of the function for other regions of Spain cannot be accepted without the necessary precautions.

The youngest 25 to 34 years old group has been eliminated from the tables, although it appears in the classical Framingham tables as the 30 years center. This age group presents a very small number of AMI cases in our country, and accordingly, the estimators were unreliable. Risk in 65 to 74 year old females for all the risk factor combinations is practically identical than in 55 to 64 year old females. This result was obtained because the quadratic factor includes the hazards function for females, and could have a physiopathological explanation, given the higher life expectancy of women, and that ischemic heart disease increases after menopause, at around 55 years of age.

The result is possibly affected by properties of the age variable coefficient for females, that was originally estimated for regions with high incidence of ischemic heart disease (Framingham). The difference between both age groups could be smaller in our environment.

The proposed correction, based on HDL-C plasma levels, simplifies the potential use of the tables. A checkbox will show a 50% increased risk if plasma levels are under 35 mg/dL, or half the risk if levels are above 59 mg/dL. The risk for individuals with 35 to 59 mg/dL plasma levels is indicated on each risk factor combination checkbox. At levels near 35 mg/dL, the risk is slightly higher (around 2%) and slightly lower at levels near 59 mg/dL (also 2%). The possibility of including the HDL-C effect in risk estimations is useful but only approximate. Nevertheless, many cases confirmed that deviations from the proposed rule rarely exceeded 3% in patients with levels below 35 mg/dL and above 60 mg/dL.

The main advantage of the Framingham equation published in 1998 is using its arterial hypertension criteria adaptation to the present criteria of the World Health Organization, categorized in 5 groups (optimal, normal, normal high, stage I and II to IV) and based on SBP and DBP values. To this regard, ischemic heart disease risks with optimal and normal blood pressure show practically no differences, as illustrated in the tables.

The tables presented are based on a calibration not validated with the population base procedure, as with hard events, or using any prospective procedure. The calibration procedure validity is reasonably guaranteed in itself.
Our population data cannot confirm that silent AMI and angina proportions, over total coronary event incidence, are similar for Girona and Framingham. Any approximations should be conservative, as a higher proportion in Girona than in Framingham is improbable.

A limitation of these predictive functions is not including some risk factors recently described (particularly, C-reactive protein, but also homocisteine, (a) lipoprotein, Chlamydia pneumoniae antibodies or Cytomegalovirus, fibrinogen and other coagulation factors and inflammation markers). Another limitation is not including any risk factor control variables, such as antihypertensive therapy or treatment of hypercholesterolemia.

The calibrated Framingham function for our environment can be a reasonably valid tool for estimating more precisely the total coronary risk in primary prevention. It should be used with an appropriate validation, not hindering any progress in elaborating specific functions for Spain using representative population cohorts with a sufficiently long follow-up period.

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REFERENCES

APPENDIX

Utilization of the predictive function is based on event probability calculation with the Cox proportional hazards method:

\[ P_{X_i(t)} = 1 - S_{X_i(t)} = 1 - e^{[-H_{0(t)} e^{(\Sigma(\beta_i x_i) - \Sigma(\beta_i x_{im})}]]} = 1 - e^{[-\ln(S_{0(t)}) e^{(\Sigma(\beta_i x_i) - \Sigma(\beta_i x_{im})}]]} = 1 - S_{0(t)} e^{(\Sigma(\beta_i x_i - \Sigma(\beta_i x_{im}))}\cdot t} \]

where \( P_{X_i(t)} \) is the coronary event probability in \( t \) time (10 years, in this case) for a patient with a group of risk factors \( x_i \) in an environment with an existing \( H_{0(t)} \) baseline event cumulative rate during \( t \) time, and with a free of coronary events \( S_{0(t)} \) baseline probability during \( t \) time; where \( \Sigma(\beta_i x_{im}) \) is an average risk linear function of the group of \( x_{im} \) values for each degree of every risk in the population; and \( \Sigma(\beta_i x_i) \) is the linear function calculated with the group of \( x_i \) values that represents each risk factor value in individuals. In both linear functions, \( \beta \) are the Cox proportional hazards function coefficients for each degree of every risk considered. \( S_{0(t)} \) is the probability of no coronary events in a patient with a group of risk factors \( x_i \) and the same conditions described; and \( e \) is the base of natural logarithms.

The estimation consists in calculating the \( \Sigma(\beta_i x_i) \) summatory by replacing \( \beta \) with the model coefficients shown in Table 1, and \( x_i \) with values of each risk factor of the patient, using «1» if the factor degree is present or «0» for the remaining degrees. As for age, \( x_i \) is replaced by years of age, and diabetes and smoking are replaced by 0 or 1, depending on their presence or not.

For example, the value for a 56 year old non-diabetic smoking male with total cholesterol 256 mg/dL, HDL-C 37 mg/dL, SBP of 146 mm Hg and DBP of 92 mm Hg, should be calculated as described in Table 3.
In the equation described above, \( H_{0(t)} \) should be calculated as:

\[
H_{0(t)} = -\ln(S_{0(t)})
\]

where \( S_{0(t)} \) is the free of events probability (expressed from 0 to 1) in males of our environment during \( t \) time=10 years

\[
H_{0(t)} = -\ln(0.951) = 0.050
\]

The difference between patient risk and average risk in the population of reference (Girona, in our case), is obtained subtracting the \( \Sigma(\beta_i x_i) \) summatory corresponding to average in males (3.489) from the \( \Sigma(\beta_i x_i) \) summatory corresponding to the patient (4.496):

\[
P_{X_i(t)} = 1 - S_{x_i(t)} = 1 - e^{-[-\ln(0.951) - (4.496 - 3.489)]} = 1 - e^{-0.050} = 0.128 = 12.8\%
\]

The probability of developing a coronary event in our environment during the next 10 years for a male characterized as mentions above is 12.8%, very similar to the probability annotated in the corresponding checkbox of the risk tables presented (12%).